

Regional economic growth and unemployment in the European Union – a spatio-temporal analysis at the NUTS-2 level (2013–2019)

MATEUSZ JANKIEWICZ¹

Abstract

The study aims to verify the relationship between the unemployment rate and economic growth in European Union (EU) regions. As the most important macroeconomic relationship, the significance of the dependence between the labour market situation and the output growth is widely known and considered. Analysis in this research was conducted using data for 229 EU regions on the NUTS-2 level in the years 2013–2019. In order to verify the relationship between the unemployment rate and the output growth, the spatio-temporal models for pooled time series and cross-sectional data (TSCS) were estimated. The Fitted Trend and Elasticity Method of verifying Okun's law was used in the analysis, wherein the deterministic trend factor was enriched with the spatial element. Educational attainment as the additional explanatory variable was included in the models. The neighbourhood between regions was quantified based on two criteria: (1) common border criterion – related to the possibility of population migrations, and (2) similarity of the unemployment rate criterion – related to the imitation effect in the issue of introduced rules and regulations on the labour market by regional governments. One of the hypotheses verified in the investigation is the superiority of the economic neighbourhood over the geographical neighbourhood.

Keywords: economic growth, European Union, Okun's law, spatio-temporal models, unemployment rate

Received August 2022, accepted April 2023.

Introduction

The problem of unemployment is one of the essential issues in macroeconomic analyses. This is widely known that the persistent regional unemployment disparities in the European Union occur (PATUELLI, R. *et al.* 2012; HALLECK VEGA, S. and ELHORST, J.P. 2016) Increasing the number of unemployed persons is a problem in both developing and developed regions. A lot of determinants significantly influence the regional unemployment disparities. For example, regional demand and supply factors, the labour migration (ANDREWS, R. 2015; LADOS, G. and HEGEDŰS, G. 2016) and amenities are the most important factors. Positive changes in these factors can disincentive to migration, compensating for relatively high unemploy-

ment rates (Rios, V. 2017). Moreover, the institutional decisions providing restrictions or incentives influence the individual decisions regarding labour demand, supply and wages paid, which changes the level of the unemployment rate (BOERI, T. 2011). Capital inadequacy can influence the decrease in the employed people above all in the developing countries. In turn, technological progress is the primary reason for the increased unemployment rate in developed regions (SOYLU, Ö.B. *et al.* 2018). Moreover, the national labour market regulation and labour market institutional system play significant role in the creation of unemployment rate in every economy. Whereas, as the most important determinant of unemployment, economic growth is pointed out. Arthur M. OKUN proved the negative relationship between

¹Department of Applied Informatics and Mathematics in Economics, Faculty of Economic Sciences and Management, Nicolaus Copernicus University in Toruń. Ul. Gagarina 13a, 87-100 Toruń, Poland. E-mail: m.jankiewicz@umk.pl – ORCID: 0000-0002-4713-778X

the unemployment rate and Gross National Product (GNP) based on the data for the United States (US) in the years 1947–1960 (OKUN, A.M. 1963). He concluded that each 1 percent increase in GNP led to a 0.3 percent decrease in the unemployment rate.

Since that time, the relationship between the unemployment rate and economic growth has been widely concerned in macroeconomic analyses. This negative short-run dependence is known as Okun's law. Previous studies conclude that the relationship formulated by OKUN is stable in many countries (BALL, L. *et al.* 2017) and possible instability is visible in terms of the economic slowdown (CAZES, S. *et al.* 2011). In turn, in some studies, authors concluded that the discrepancy in Okun's relationship between regions within one country occurs (ADANU, K. 2005; BINET, M. and FACCHINI, F. 2013; DURECH, R. *et al.* 2014).

There are three methods of verifying Okun's relationship: (1) Trial Gaps Method, (2) First Differences Method, and (3) Fitted Trend and Elasticity Method (BARRETO, H. and HOWLAND, F. 1993). The first two methods treated analysed processes as the processes stationary in the variance. Instead, within the meaning of the third method, processes are stationary in the average. The First Differences Method is characterized by the greatest popularity in previous studies. In this research, the Fitted Trend and Elasticity method is used, but the deterministic trend factor is enriched with the spatial element (originally, only time tendency was considered).

In this study, the relationship between the unemployment rate and its main determinants – economic growth and educational attainment – is analysed. The study's main aim is to show that the economic growth and unemployment rate are significantly related in European Union regions. Moreover, the stronger importance of economic similarity between territorial units than their near geographical location in the considered relationship is verified. As a space and time range of the research, the NUTS-2 European Union regions in the years 2013–2019 were chosen

(due to lack of data for the unemployment rate, Croatian regions were omitted). In the verification of the mentioned relationship, the spatial and spatio-temporal dependencies were included. Many researchers pointed out the importance of spatial connections in the unemployment rate analyses (OVERMAN, H. *et al.* 2002; PATACCHINI, E. and ZENOU, Y. 2007; HALLECK VEGA, S. and ELHORST, J.P. 2016). In this study, two types of neighbourhood connections were considered. The first is the geographical neighbourhood (associated with the possibility of migration), and the second is the economic neighbourhood (associated with the unemployment rate similarity). Two research hypotheses were verified in this investigation: (1) Output growth and education have a significant positive impact on the labour market conditions in the EU regions, and (2) Economic similarity of regions is more important than a geographical neighbourhood in the formation of Okun's relationship.

There are many studies considering Okun's relationship at the regional level, and different methods are used in order to verify it. A couple of studies pertain to Okun's relationship in the Spanish provinces (VILLAVERDE, J. and MAZA, A. 2007, 2009; CLAR-LOPEZ, M. *et al.* 2014; CHÁFER, C.M. 2015; BANDE, R. and MARTÍN-ROMÁN, Á. 2018; GUISINGER, A.Y. *et al.* 2018; CUTANDA, A. 2023). All of these investigations are based on non-spatial analysis, and only VILLAVERDE, J. and MAZA, A. (2007) considered the Fitted Trend and Elasticity Method with the quadratic trend. BANDE, R. and MARTÍN-ROMÁN, Á. (2018) estimated a simple model for the first differences of processes and also for trial gaps. The same method of research took CLAR-LOPEZ, M. *et al.* (2014). The subject of the investigations in terms of the relationship between unemployment rates were also regions from other European countries, e.g., Italian provinces (SALVATI, L. 2015), Finnish regions (KANGASHARJU, A. *et al.* 2012), Greek regions (APERGIS, N. and REZITIS, A. 2003), and also Czech and Slovak regions (DURECH, R. *et al.* 2014). Most of these studies underline the regional disparities in the unemploy-

ment rate formation rely on their economic development. In turn, YERDELEN, F. and İÇEN, H. verified Okun’s relationship for NUTS-2 level regions from 20 European countries using panel data models (YERDELEN, F. and İÇEN, H. 2019). Apart from the studies for European countries, it is possible to find analyses of the mentioned relationship at the regional level in the United States (HUANG, H.C. and YEH, C.C. 2013; GUISSINGER, A.Y. *et al.* 2018), Canada (ADANU, K. 2005), Indonesia (SASONKO, G. *et al.* 2020), and South Africa (KAVESE, K. and PHIRI, A. 2020).

The spatial factor in the analysis of the dependence between the unemployment rate and output growth is also included in previous studies. DURAN, H.E. considered this dependence using spatial panel data models using data for 26 Turkish NUTS-2 regions (DURAN, H.E. 2022). Spatial regression models were used in the analyses concerning the unemployment rate in the United States (MONTERO-KUSCEVIC, C.M. 2011; PERREIRA, R.M. 2013), and EU15 NUTS-2 regions (HERWARTZ, H. and NIEBUHR, A. 2011). In turn, Adolfo MAZA conducted the analysis for the widest space range in the mentioned relationship (MAZA, A. 2022). He considered the unemployment rate in 265 European regions between 2000 and 2019.

Methodology

In this research, spatial econometric methods were used in order to verify the Okun’s relationship. Spatial econometric models contain the influence of process changes in the neighbouring regions on the same process in the established region. In the regional analyses, connections between nearby and also similar (in the economic context) units are very important, in particular in case of the unemployment analysis. The economic conditions and possibilities of the neighbours can encourage people to jobs migrations, changing the labour situation in the considered unit.

In the first part of the investigation, the spatio-temporal structure of processes was

analysed. The structure is composed of the spatio-temporal trend and spatio-temporal autocorrelation. Initially, spatio-temporal trend models were considered, which general form is as follows (CRESSIE, N.A.C. 1993):

$$P(s_i, t) = \sum_{k=0}^p \sum_{m=0}^p \sum_{l=0}^p \theta_{kml} x_i^k y_i^m t^l, \tag{1}$$

where $s_i = [x_i, y_i]$ denotes unit’s location coordinates on the plane (longitude and latitude, respectively), $i = 1, 2, \dots, N$ are indexes of spatial units, and p means the polynomial trend degree ($k + m + l \leq p$) but t indicates time.

Simultaneously, the spatio-temporal autocorrelation presence as the second element of the spatio-temporal structure was checked. The spatial autocorrelation is tested using Moran statistics, which takes the following form (MORAN, P.A.P. 1948; SCHABENBERGER, O. and GOTWAY, C.A. 2005):

$$I = \frac{1}{\sum_{i=1}^n \sum_{j=1}^n w_{ij,t}} \cdot \frac{\sum_{i=1}^n \sum_{j=1}^n w_{ij,t} [y_{i,t} - \bar{y}][y_{j,t} - \bar{y}]}{\frac{1}{n} \sum_{i=1}^n [y_{i,t} - \bar{y}]^2} = \frac{n}{S_0} \cdot \frac{z^T W^* z}{z^T z}. \tag{2}$$

where $y_{i,t}$ is the observation of the process in the i^{th} region in time t , \bar{y} denotes the average value of the process, W^* is the block matrix of spatio-temporal connections between units given as SZULC, E. and JANKIEWICZ, M. (2018):

$$W^* = [w_{ij,t}]_{NT \times NT} = \begin{bmatrix} W_1 & 0 & \dots & 0 \\ 0 & W_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & W_T \end{bmatrix}, \tag{3}$$

wherein $W_1 = W_2 = \dots = W_T$ are standard spatial connectivity matrices quantified for a certain year. In this study, these matrices are the same for all years.

In this research, two types of row-standardized to unity matrices were adopted. The first of them is based on the common border criterion (marked as W). Therefore, two regions are neighbours if they have a common land border. In turn, the second defines the neighbourhood as the economic similarity (D) – regions are neighbours if the difference between their unemployment rate level in the last year of the investigation does not exceed a certain specific value (established as 0.8% – the 15th percentile of differences between the unemployment rate in all regions). The procedure of building the economic dis-

tance matrix is presented by JANKIEWICZ, M. and SZULC, E. in their study (JANKIEWICZ, M. and SZULC, E. 2021).

Statistically significant Moran's I coefficient signalizes the presence of spatial autocorrelation. Its positive value denotes that territorial units create clusters of the regions with a similar level of the analysed phenomenon. In turn, the negative sign of statistics points out that neighbouring regions are characterized by different values of the considered process. Non-significant statistics testifies to a random distribution of the process values in space.

Next, the spatio-temporal models of the relationship between the unemployment rate, economic growth, and educational attainment were considered. The general form of the TSCS model (pooled time series and cross-sectional data model) is as follows:

$$Y_{i,t} = \beta_1 X_{1i,t} + \beta_2 X_{2i,t} + \varepsilon_{i,t}, \quad (4)$$

where $Y_{i,t}$ denotes the unemployment rate in the i^{th} region in time t , $X_{1i,t}$ and $X_{2i,t}$ are levels of Gross Domestic Product per capita and educational attainment, respectively (all expressed in natural logarithms). In turn, $\varepsilon_{i,t}$ indicates the spatio-temporal random component, but β_1 and β_2 are the structural parameters. Logarithms of variables cause that parameter β_1 is the elasticity parameter in the Fitted Trend and Elasticity Method in the Okun's law verification. Model (4) is deprived of constant due to all considered variables are filtered out from deterministic spatio-temporal trend, which is responsible for their average values.

In terms of global spatial autocorrelation in the residuals of the model (4) the character of the spatial dependence was determined using Lagrange Multiplier (LM) tests in the basic and robust version (ANSELIN, L. *et al.* 2004). Including spatially lagged explanatory variables in the models, the spatio-temporal Durbin model (STDM) and spatio-temporal hybrid model (STHM) are given as follows:

$$Y_{i,t} = \beta_1 X_{1i,t} + \beta_2 X_{2i,t} + \rho \sum_{j \neq i} w_{ij,t} Y_{j,t} + \theta_1 \sum_{j \neq i} w_{ij,t} X_{1j,t} + \theta_2 \sum_{j \neq i} w_{ij,t} X_{2j,t} + \varepsilon_{i,t}, \quad (5)$$

$$Y_{i,t} = \beta_1 X_{1i,t} + \beta_2 X_{2i,t} + \theta_1 \sum_{j \neq i} w_{ij,t} X_{1j,t} + \theta_2 \sum_{j \neq i} w_{ij,t} X_{2j,t} + \eta_{i,t}, \quad (6)$$

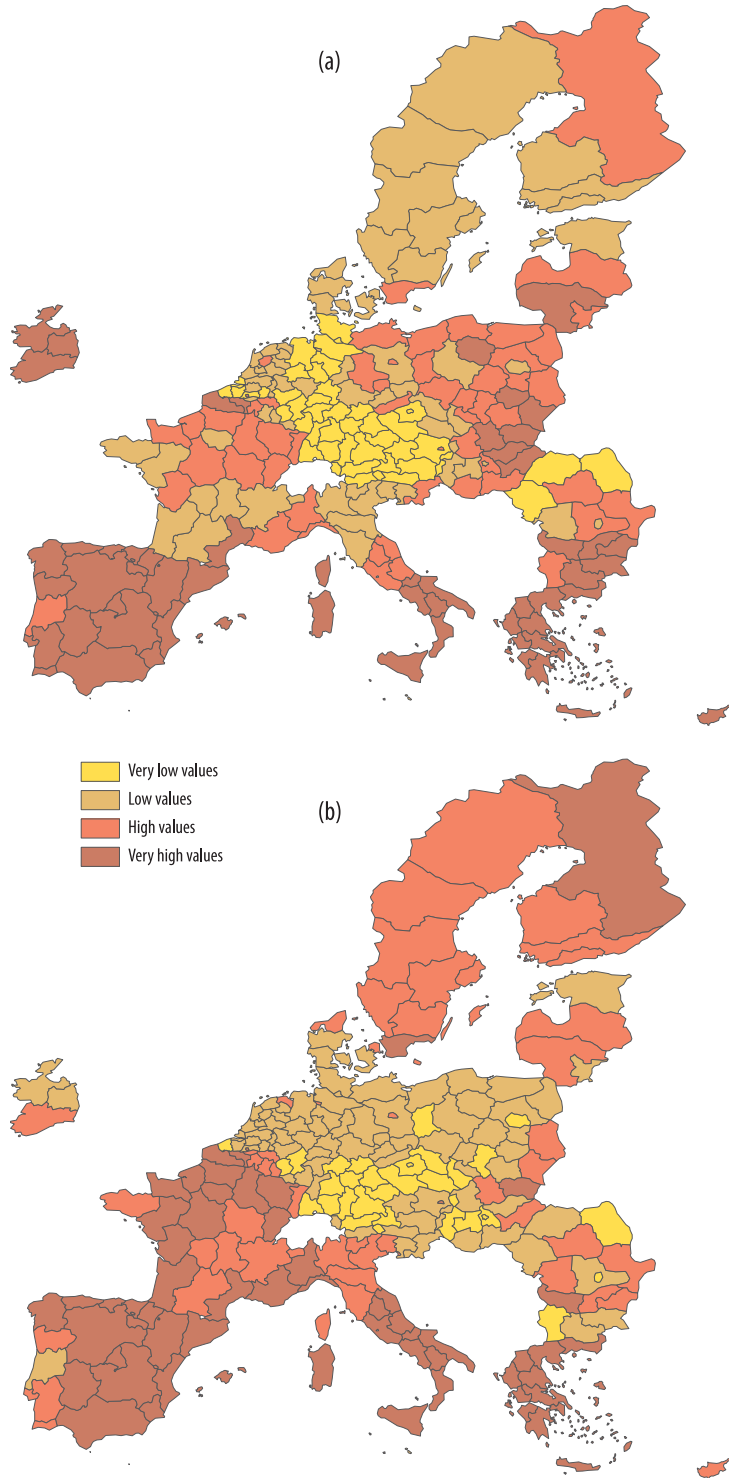
$$\eta_{i,t} = \lambda \sum_{j \neq i} w_{ij,t} \eta_{j,t} + \varepsilon_{i,t},$$

where $Y_{i,t}$, $X_{1i,t}$, $X_{2i,t}$, $\varepsilon_{i,t}$, β_1 , β_2 – as above, $\sum_{j \neq i} w_{ij,t}$, $Y_{j,t}$, $\sum_{j \neq i} w_{ij,t} X_{1j,t}$, $\sum_{j \neq i} w_{ij,t} X_{2j,t}$ – spatially lagged variables, $\sum_{j \neq i} w_{ij,t} \eta_{j,t}$ – spatially lagged random process, θ_1 , θ_2 , ρ , λ – structural parameters. Parameters ρ and λ evidence the spatial dependence between neighbouring territorial units.

Spatial and spatio-temporal structure of processes

Data used in this study concern the unemployment rate (marked as Y), Gross Domestic Product per capita (X_1), and educational attainment level, understood as the percent of the population that graduated upper secondary and post-secondary (not tertiary) school (X_2) in the European Union regions in the years 2013–2019. Better economic conditions in regions favour the creation of new workplaces, whereas the higher education level of society improves chances of getting a job. The indicators used in the research are only a few of many that significantly impact regional unemployment, but they are considered the most important. Apart from them, e.g., the innovation level and the economic structure of regions are important. The first is not included due to data unavailability for the whole considered area, while the second will be of interest for further research. Moreover, in case of unemployment, the age structure of the population and distance from main urban centres play an important role. The established period is the maximum period that can be analysed in light of the data availability. Moreover, due to a lack of data characterizing the unemployment level, Croatia's regions were omitted. All data come directly from the European Statistical Office (EUROSTAT) database – <https://ec.europa.eu/eurostat/data/database> (accessed: 04.07.2022).

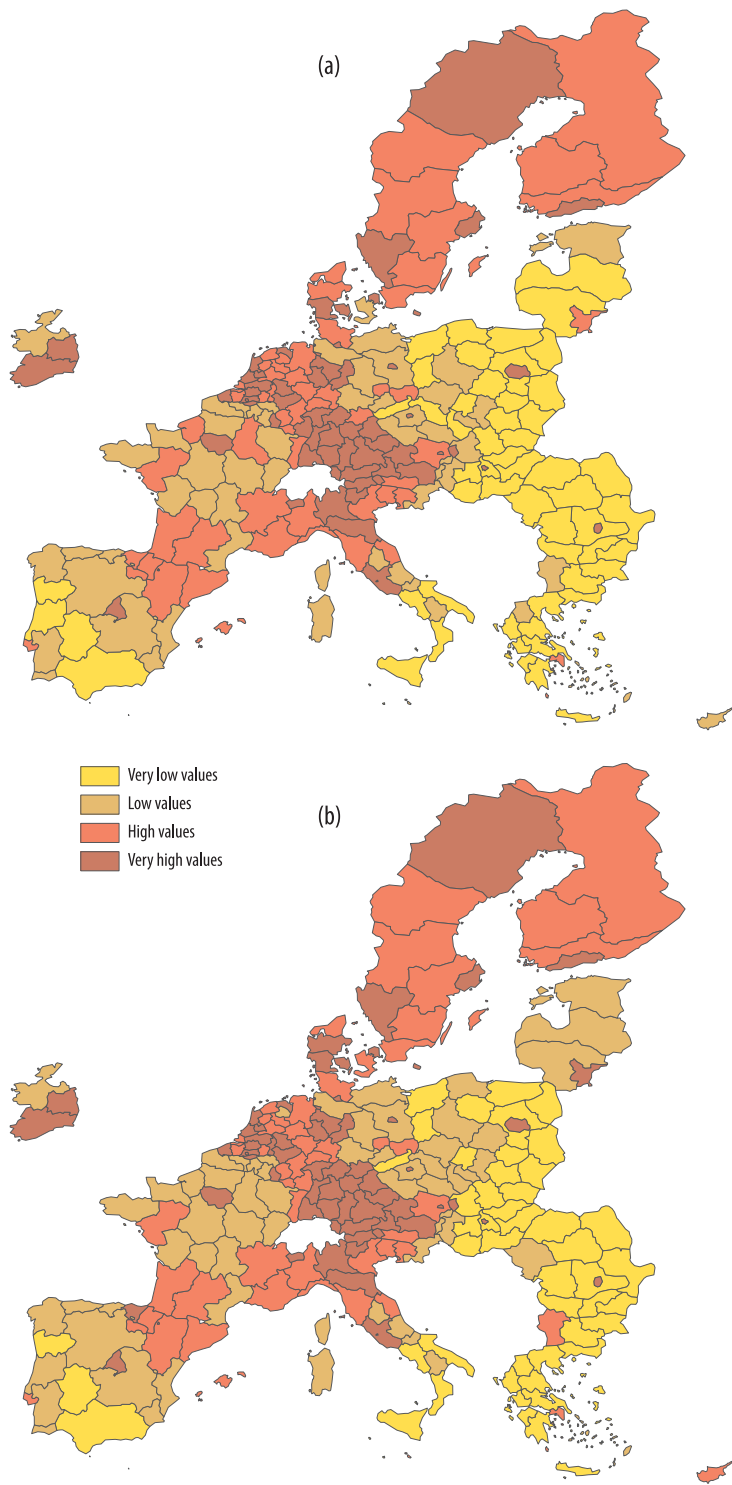
In the first part of the research, spatial distributions of considered processes were presented. These distributions were shown in three figures (for variable Y , for variable X_1 and for



variable X_2) in the extreme years of the analysis. In each figure, part (a) indicates the distribution in 2013, but part (b) refers to spatial differentiation in 2019. EU regions were divided into four groups using positional measures of the descriptive statistics (median and quarter deviation).

As we can see in *Figure 1*, the highest unemployment rate in 2013 was observed in the Iberian Peninsula regions, Greek regions, and the units located in Southern Italy. Moreover, relatively high unemployment was noted in most of Eastern Europe NUTS-2 level regions (above all in the regions located in Lithuania, Poland, and the Slovak Republic). On the other hand, Austrian and West German regions were characterized by the best labour market conditions – the unemployment rate was relatively low. This is worth noting the low level of the considered variable in the North Romanian regions. In 2019 the situation in the labour market in the EU was slightly different than in 2013. The unemployment level in all Francian units was above the median in the last year of the investigation. Instead, the

Fig. 1. Spatial distribution of the unemployment rate in EU regions in the years 2013 (a) and 2019 (b)

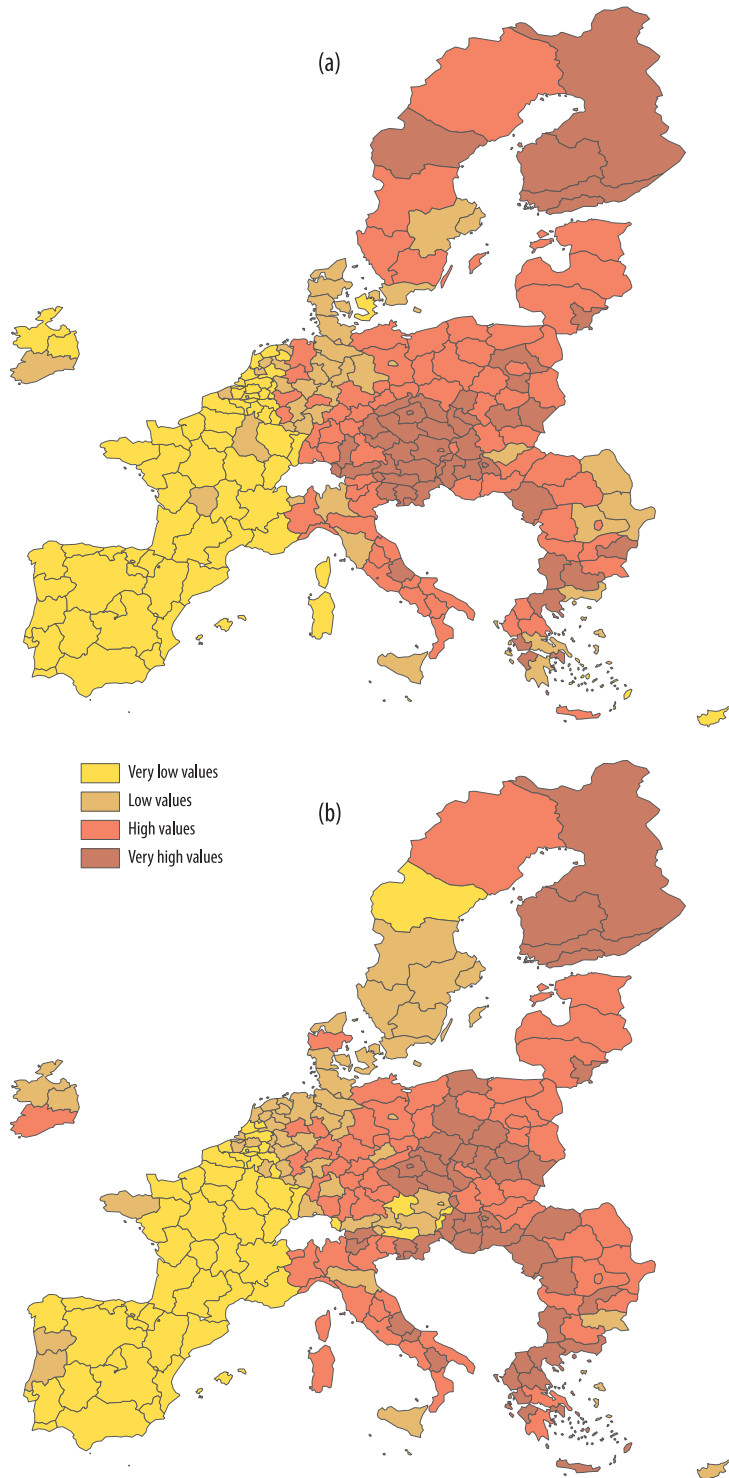


variable values in 2013 were more diversified between these regions. The relative deterioration of the situation concerns North Italian provinces and Scandinavian regions as well. In contrast, most of the Polish provinces found themselves in units with unemployment rate values below the median, which denotes the relative improvement in this part of the European Community.

Based on spatial distributions of the unemployment level in the EU regions, it can be presumed that values of the considered variable exhibit a certain tendency in space. Therefore, the spatial factor in the analysis should be included. In this connection, the spatio-temporal trend models in the following part of the study were concerned.

Seeing the spatial distributions of GDP per capita (Figure 2), we can note that the spatial differentiation of process values in both analysed years was analogous. Relatively high economic growth (with the values of GDP per capita above median) was observed in the Central EU regions (located in Austria,

Fig. 2. Spatial distribution of the GDP per capita (PPS) in EU regions in the years 2013 (a) and 2019 (b)



West Germany, North Italy, and Benelux countries). Also, a high GDP level was noted in the Scandinavian and South Ireland regions. On the other hand, the relatively less developed units were located above all in the eastern part of the European Community, except for regions Bratislava (SK01), Praha (CZ01), Warszawski stołeczny (PL91), Sostinès (LT01), and București-Ifov (RO32). Three mentioned units belonged to the group of regions with the highest values of the GDP per capita in 2019. A similar situation was observed in the Iberian Peninsula provinces. Almost all regions were classified into the groups of very low and low economic growth levels. Only two regions were characterized by economic growth above the median – one Spanish and one Portuguese (Madrid – ES30 and Área Metropolitana de Lisboa – PT17, respectively). As in the case of the unemployment rate, certain spatial tendencies in the formation of GDP per capita values were observed.

Figure 3 presents spatial distributions of educational attainment.

Fig. 3. Spatial distribution of the educational attainment level in EU regions in the years 2013 (a) and 2019 (b)

The percent of graduates in the upper secondary and post-secondary school (excluding tertiary education) is considered. It is worth seeing the possibility of the division of EU regions into two parts in both extreme years of the study (more visible in 2019). The first was in the eastern part of the European Community, where almost all regions were characterized by a relatively high percentage of graduates in upper secondary and post-secondary school. Instead, the western part of the mentioned area was dominated by units with the values of the considered process below the median. It is worth noting the lowest level of the variable X_2 in the French and Spanish regions, where the unemployment rate was relatively high. Units with the low and high values of the educational attainment process created two almost coherent areas, which lead to presumption about a certain spatial tendency in their formation.

The observations made based on the spatial distributions of all processes allowed us to consider the two-dimensional deterministic trend (with the spatial and time factors) in order to filter out long-term tendencies. *Table 1* shows the results of estimation and verification of the spatio-temporal trend models for all variables. In the models, only the statistically significant parameters were left.

It is a difference in the degree of trend obtained for the variable Y and the two remain-

ing variables. The unemployment rate in the period 2013–2019 was shaped according to the second-degree spatio-temporal trend. Considering estimates of parameters θ_{100} and θ_{010} for variables X_1 and X_2 , we can conclude that their values averagely have been growing in the western-northern and eastern-northern directions, respectively. This confirms the insights visible in figures 2 and 3. Moreover, positive estimates of parameter θ_{001} indicate the average increase of the GDP per capita and educational attainment in the years 2013–2019. In addition, an average decrease of the variable Y towards the north and east should be noted (negative estimations of parameters θ_{010} and θ_{100} respectively).

The low level of the determination coefficient R^2 is the characteristic feature of the spatial and spatio-temporal trend models. However, we can see that the least coherent values formation in space and time was pertaining to Gross Domestic Product per capita.

Additionally, the Moran test evaluating the dependence between neighbouring regions was conducted. In the spatial autocorrelation analysis, two types of neighbourhood matrices were used. First of them defines neighbouring regions as regions with a common land border (geographical neighbourhood – W). Instead, the second points out that two units are neighbours if they had a similar level of the unemployment rate in 2019 (eco-

Table 1. The results of estimation and verification of the spatio-temporal trend models

Parameter	Unemployment (Y)		GDP per capita (X_1)		Educational attainment (X_2)	
	Estimate	p-value	Estimate	p-value	Estimate	p-value
θ_{000}	15.6290	0.0000	8.9967	0.0000	3.8209	0.0000
θ_{100}	-0.0185	0.0000	-0.0190	0.0000	0.0120	0.0000
θ_{010}	-0.5110	0.0000	0.0259	0.0000	0.0049	0.0000
θ_{001}	–	–	0.0301	0.0000	0.0023	0.0927
θ_{200}	0.0008	0.0000	–	–	–	–
θ_{020}	0.0048	0.0000				
θ_{101}	-0.0014	0.0123				
θ_{002}	-0.0104	0.0000				
R^2	0.5432		0.3315		0.5318	
Moran test						
Matrix	I	p-value	I	p-value	I	p-value
W	0.4714	0.0000	0.3649	0.0000	0.4220	0.0000
D	0.4953	0.0000	0.0466	0.0019	0.1096	0.0000

conomic distance matrix – D). The Moran’s I statistics for variable Y are very similar using both types of the connection matrix (0.4714 and 0.4953, respectively) with the proviso that units with a similar unemployment rate in 2019 showed slightly higher neighbourhood dependence. In contrast, the Moran’s I coefficient evaluated considering the distance matrix was relevantly lower than in the case of the first-order contiguity matrix W) for two remaining variables. It means that the dependence of GDP per capita and educational attainment between neighbouring regions in a geographical space was stronger than between regions neighbouring in economic terms. Nonetheless, all determined Moran’s coefficients turned out statistically significant. This situation indicates the necessity of spatial dependence inclusion in the analysis of the relationship between the unemployment rate and economic growth, and educational attainment.

Okun’s relationship models

In the first part of the relationship analysis, the TSCS (pooled time series and cross-sectional) model was considered. *Table 2* presents the results of the estimation and verification of the pooled spatio-temporal model. This is Okun’s model extended with the educational factor.

The negative estimate of statistically significant parameter β_1 denotes that an increase in the GDP per capita causes an average decrease in the unemployment rate. This confirms the observations made by Arthur M. OKUN. The value of the elasticity parameter indicates that an increase in the GDP by 1 percent provides to decrease in the unemployment rate averagely by 0.41 percent *ceteris paribus*. A higher strength of the influence shows the educational level, where an increase of 1 percent causes the average decrease in the unemployment rate averagely by 0.73 percent.

Moran test results indicate the presence of global spatio-temporal autocorrelation in the model residuals. Therefore, the significance of dependence between neighbouring regions in the light of both neighbourhood matrices was concluded. So the spatio-temporal models estimated considering geographical and economic neighbourhoods should be analysed.

In order to determine the character of spatio-temporal dependence, the Lagrange Multiplier tests were conducted. Tests statistics in the basic version (LM_{err} and LM_{lag}) do not solve the problem of the model choice – both are statistically significant. Analysing LM statistics in the robust versions of tests, we can conclude that the model with spatial factor in the error term for the W matrix is better. In turn, for the economic distance matrix (marked as D), the model with a spatial lag of the dependent variable should be

Table 2. The results of estimation and verification of the TSCS extended Okun’s model

Parameter	Estimate	Standard error	t statistics	p-value
β_1	-0.4194	0.0320	-13.1190	0.0000
β_2	-0.7301	0.0925	-7.8960	0.0000
R_2	0.1353			
Moran test				
Matrix	W		D	
I	0.4363		0.4257	
p-value	0.0000		0.0000	
LM tests				
Statistics	Estimate	p-value	Estimate	p-value
LM_{err}	611.5708	0.0000	699.1660	0.0000
LM_{lag}	575.7326	0.0000	859.9580	0.0000
RLM_{err}	38.3835	0.0000	21.3310	0.0000
RLM_{lag}	2.5453	0.1106	182.1230	0.0000

chosen. Regardless of the results of the robust *LM* tests, both types of spatio-temporal models enriched with the spatial lags of explanatory variables were estimated.

Table 3 presents the results of estimation and verification of the spatio-temporal Durbin models (STDM) and spatio-temporal hybrid models (STHM), considering both connection matrices.

Similar to the TSCS model, estimates of statistically significant parameters β_1 and β_2 are negative, which confirms the positive influence of the GDP per capita and educational attainment increase on the labour market conditions in the EU regions. Nonetheless, the strength of impact is lower. The highest difference is observed in the estimate of the parameter β_2 in the models for the *D* matrix – the 1 percent increase in educational attainment provides the average decrease in the unemployment rate around 0.4 percent (less than for the TSCS model around 0.3%). In all models, parameters ρ and λ are statistically significant, which confirms the necessity of including spatial factors in the analysis. Values of estimates of the spatial parameters within one connection matrix are very similar. Moreover, the estimates of both consid-

ered parameters are slightly lower in the case of models for the economic distance matrix. This indicates that regions neighbouring in the geographical space were more similar in the unemployment rate than the neighbours determined by the economic terms.

It is worth noting the statistical significance of the θ_1 parameter in all estimated models. Except for the spatial Durbin model for the *W* matrix (SDM_W), the estimate of this parameter is negative. Considering the geographical neighborhood, changes in the unemployment rate in the neighboring regions had a different impact than changes in the random processes or processes omitted in the model. In turn, shocks like an increase in the unemployment rate level or in the random processes in the neighbouring (from the economic point of view) units caused a significant decrease in the unemployment rate in a certain region. However, shocks in the random processes or omitted explanatory variables were slightly stronger.

The desirable characteristic of the models with the *W* matrix is the lack of spatial autocorrelation in the models' residuals. In the light of the Akaike Criterion (AIC) and the logarithm of likelihood values (Log-lik)

Table 3. The results of estimation and verification of the spatial extended Okun's model

Parameter	Model			
	STDM_W	STHM_W	STDM_D	STHM_D
β_1	-0.3094 (0.0000)	-0.3289 (0.0000)	-0.3111 (0.0000)	-0.3296 (0.0000)
θ_1	0.1461 (0.0014)	-0.1710 (0.0113)	-0.2275 (0.0001)	-0.3211 (0.0000)
β_2	-0.6009 (0.0000)	-0.6310 (0.0000)	-0.4030 (0.0000)	-0.4448 (0.0000)
θ_2	0.2530 (0.0761)	-0.2921 (0.1194)	-0.0718 (0.6167)	-0.2769 (0.1483)
ρ	0.6204 (0.0000)	–	0.5642 (0.0000)	–
λ	–	0.6249 (0.0000)	–	0.5822 (0.0000)
Diagnostics				
Moran test	-0.0155 (0.2127)	-0.0170 (0.1895)	-0.0490 (0.0015)	-0.0347 (0.0182)
AIC	1,126.5000	1,118.9000	1,052.4000	1,095.5000
Log-lik	-557.2697	-553.4698	-520.1802	-541.7608

the best model is the spatial Durbin model estimated using connection matrix D based on the economic distance neighbourhood. Therefore, we can conclude about higher cognitive values of the models with the economic neighbourhood (regardless of the spatial autocorrelation presence in the model residuals).

Statistical significance of the parameters ρ , β_1 and θ_1 allows for quantifying the short-term spatial spillovers. Nevertheless, this issue is not the subject of the study. The quantification of the spatial effects is one of the directions of further analysis.

Discussion

The analysis presented in this paper shows a large variation in the unemployment rate between EU regions, which is a significant economic problem. The problem particularly refers to the southern and eastern parts of the European Union. The causes of this situation have a different character. The economic growth and education level of society are considered as the most important indicators influencing the labour market level. Estimated models confirm that the increase in Gross Domestic Product per capita and educational attainment (measured by the percentage of graduates in upper secondary and post-secondary school) significantly causes the decrease in the unemployment rate. Moreover, the other individual characteristics of regions influence the labour market situation. One of them is specific economic structure, i.e., if the considered regions are rural or industrial. A verification of Okun's relationship for EU regions taking into account their specificity will be the subject of further research. BOĐA, M. and POVAŽANOVÁ, M. (2020) point out the necessity of diversifying regions by their specific characters estimating Okun's relationship. In turn, BONAVENTURA, L. *et al.* (2018) verified this relationship in two gender groups in Italy. They inferred differences in the sensitivity level of the unemployment rate on changes in GDP per capita between males and females depending on the geographic location of the region.

Some of the authors so far analysed Okun's relationship on the regional level. For example, DURAN, H.E. (2022) showed its significance for Turkish provinces. He did not include additional explanatory variables apart from the economic growth level, the increase of which causes the decrease in the unemployment rate in all provinces. In turn, MELGUIZO, C. (2017) considered the connection between economic growth and the unemployment rate in Spanish provinces. She inferred the same type of relationship (with different strengths) throughout the area. Also negative sign of Okun's coefficient for all regions of Slovenia obtained DAJCMAN, S. (2018). PALOMBI, S. *et al.* (2017) showed the same for Great Britain, analysing data for regions at a NUTS-3 level. Their study is one of few using spatial econometric models as a research tool for the verification of Okun's relationship. Other analyses based on the spatial econometric approach were conducted by MONTERO-KUSCEVIC, C.M. (2011), FERREIRA, R.M. (2013), and MAZA, A. (2022). In this research, additionally, the educational attainment level was included in Okun's model, which is not found in many other studies.

The methodological approach used in this research differs from other approaches. Firstly, previous researches including spatial connections between territorial units based on the First Differences Method of Okun's law. This analysis used the Fitted Trend and Elasticity Method enriched with the spatial trend. It is a new approach to establishing long-term dependencies in the formation of key indicators used in the investigation, which treats the trend wider than yet. In turn, the definition of one of the spatial connection matrices is new. It is an economic distance matrix built on the unemployment rate similarity between regions. As we saw in the results, changes in GDP per capita level in regions with similar unemployment rates influence stronger on the labour market conditions in the specific region, than changes in the regions directly adjacent. Spatial models in previous studies were estimated using neighbourhood matrices built based on the common land border or the geographic distance criteria. The weakness of this research is not consider-

ing the specific characteristics of regions, for example, a population composition, an economic structure, and other important indicators, such as an innovation level. These aspects will be the subject of further research.

Conclusions

The regional approach to the verify of Okun's relationship has become more and more popular in macroeconomic analyses. Regardless of the regional disparities in the unemployment rate and the economic growth between NUTS-2 level units, the general dependence among these processes was confirmed in the European Union. The Okun's elasticity parameter (β_1) in the estimated spatial models took a value similar to this, considered to be a benchmark around -0.3 (value received by Arthur M. OKUN in his study). Moreover, the educational attainment turned out to be significant, and an increase in the percentage of graduates in upper secondary and post-secondary schools caused a decrease in the unemployment rate. In this connection, the first research hypothesis of the study was confirmed.

The economic similarity included in the models in the form of a neighbourhood matrix turned out to be statistically significant, so the similarity between regions related to the unemployment rate is relevant in Okun's relationship verification. So this is the second type of connection, next to the repeatedly confirmed significance of the geographical closeness (in this study, too), which allows for an understanding of the formation of the relationship between unemployment and output growth. A comparison of the estimated spatio-temporal models shows that models with the economic neighbourhood (regardless of the certain imperfections) better explain the mentioned dependence, which confirms the second research hypothesis. It means that the regions similar in the unemployment rate levels are connected stronger in case of the relationship between economic growth and labour market conditions than the regions directly adjacent to each other.

The proximity in the sense of the similarity of the unemployment rate can explain the imitation effect related to regularities and rules introduced by the governments of regions. The patterning of the regional rulers' behaviours from other provinces in the case of the labour market situation can provide similar changes in the labour market in a certain unit. It is also worth noting the policy of combating unemployment should be fitted to the regional specificity of the local labour market.

It is worth noting that in the adopted time range (2013–2019), all crises are omitted: (1) financial crisis in 2007–2009, (2) economic slowdown in 2012, and (3) COVID-19 pandemic from 2020 (due to lack of the data). In this connection, the relationship between the unemployment rate and output growth may be accepted as relatively stable in the European Union regions (which does not mean that not differentiated between units). In further research, this is worth concerning also the analysis of the regimes of the regions divided by the economic growth level and the impact of the COVID-19 pandemic on the mentioned relationship. This research will be enriched with the spatial effects quantified based on the estimated models and the use of other spatial connections matrices.

REFERENCES

- ADANU, K. 2005. A cross-province comparison of Okun's coefficient for Canada. *Applied Economics* 37. (5): 561–570. Doi: 10.1080/0003684042000201848
- ANDREWS, R. 2015. Labour migration, communities and perceptions of social cohesion in England. *European Urban and Regional Studies* 22. (1): 77–91. Doi: 10.1177/0969776412457165
- ANSELIN, L., FLORAX, R. and REY, S.J. 2004. *Advances in Spatial Econometrics. Methodology, Tools and Applications*. New York, Springer Verlag.
- APERGIS, N. and REZITIS, A. 2003. An examination of Okun's law: Evidence from regional areas in Greece. *Applied Economics* 35. (10): 1147–1151. Doi: 10.1080/0003684032000066787
- BALL, L., LEIGH, D. and LOUNGANI, P. 2017. Okun's law: Fit at 50? *Journal of Money, Credit and Banking* 49. (7): 1413–1441. Doi: 10.1111/jmcb.12420
- BANDE, R. and MARTÍN-ROMÁN, Á. 2018. Regional differences in the Okun's relationship: New evidence for

- Spain (1980–2015). *Investigaciones Regionales – Journal of Regional Research* 41. 137–165.
- BARRETO, H. and HOWLAND, F. 1993. *There Are Two Okun's Law Relationships between Output and Unemployment*. Crawfordsville, Wabash College.
- BINET, M. and FACCHINI, F. 2013. Okun's law in the French regions: A cross-regional comparison. *Economics Bulletin* 33. (1): 420–433.
- BOĀA, M. and POVAŽANOVÁ, M. 2020. Formal and statistical aspects of estimating Okun's law at a regional level. *Papers in Regional Science* 99. (4): 1113–1136. Doi: 10.1111/pirs.12511
- BOERI, T. 2011. Institutional reforms and dualism in European labour markets. In *Handbook of Labour Economics* Vol. 4b. Eds.: ASHENFELTER, O. and CARD, D., Elsevier Publication, 1173–1236.
- BONAVENTURA, L., CELLINI, R. and SAMBATATO, M. 2018. *Gender Differences in Okun's Law across the Italian Regions*. Munich, MPRA Paper No. 87557. Available at https://mpra.ub.uni-muenchen.de/87557/1/MPRA_paper_87557.pdf
- CAZES, S., VERICK, S. and AL HUSSAMI, F. 2011. *Diverging Trends in Unemployment in the United States and Europe: Evidence from Okun's Law and the Global Financial Crisis*. Employment Sector, Employment Working Paper No. 106. Geneva, ILO.
- CHÁFER, C.M. 2015. *An Analysis of the Okun's Law for the Spanish Provinces*. Working Paper 2015/01. Barcelona, Institut de Recerca en Economia Aplicada Regional i Pública.
- CLAR-LOPEZ, M., LÓPEZ-TAMAYO, J. and RAMOS, R. 2014. Unemployment forecasts, time varying coefficient models and the Okun's law in Spanish regions. *Economics and Business Letters* 3. (4): 247–262.
- CRESSIE, N.A.C. 1993. *Statistics for Spatial Data*. New York, John Wiley & Sons Inc.
- CUTANDA, A. 2023. Stability and asymmetry in Okun's law: Evidence from Spanish regional data. *Panoeconomicus* 70. (2): 219–238. Doi: 10.2298/PAN191203012C
- DAJCMAN, S. 2018. A regional panel approach to testing the validity of Okun's law: The case of Slovenia. *Economic Computation & Economic Cybernetics Studies & Research* 52. (3): 39–54. Doi: 10.24818/18423264/52.3.18.03
- DURAN, H.E. 2022. Validity of Okun's law in a spatially dependent and cyclical asymmetric context. *Panoeconomicus* 69. (3): 447–480. Doi: 10.2298/PAN190529003D
- DURECH, R., MINEA, A., MUSTEA, L. and SLUSNA, L. 2014. Regional evidence on Okun's law in Czech Republic and Slovakia. *Economic Modelling* 42. (C): 57–65. Doi: 10.1016/j.econmod.2014.05.039
- GUISINGER, A.Y., HERNÁNDEZ-MURILLO, R., OWYANG, M.T. and SINCLAIR, T.M. 2018. A state-level analysis of Okun's law. *Regional Science and Urban Economics* 68. (C): 239–248. Doi: 10.1016/j.regsciurbeco.2017.11.005
- HALLECK VEGA, S. and ELHORST, J.P. 2016. A regional unemployment model simultaneously accounting for serial dynamics, spatial dependence and common factors. *Regional Science and Urban Economics* 60. (C): 85–95. Doi: 10.1016/j.regsciurbeco.2016.07.002
- HERWARTZ, H. and NIEBUHR, A. 2011. Growth, unemployment and labour market institutions: evidence from a cross-section of EU regions. *Applied Economics* 43. (30): 4663–4676. Available at <https://www.tandfonline.com/doi/full/10.1080/00036846.2010.493142>
- HUANG, H.C. and YEH, C.C. 2013. Okun's law in panels of countries and states. *Applied Economics* 45. (2): 191–199. Doi: 10.1080/00036846.2011.597725
- JANKIEWICZ, M. and SZULC, E. 2021. Analysis of spatial effects in the relationship between CO₂ emissions and renewable energy consumption in the context of economic growth. *Energies* 14. (18): 5829. Doi: 10.3390/en14185829
- KANGASHARJU, A., TAVERA, C. and NIJKAMP, P. 2012. Regional growth and unemployment: The validity of Okun's law for the Finnish regions. *Spatial Economic Analysis* 7. (3): 381–395. Doi: 10.1080/17421772.2012.694141
- KAVESE, K. and PHIRI, A. 2020. A provincial perspective of nonlinear Okun's law for emerging markets: The case of South Africa. *Studia Universitatis „Vasile Goldis” Arad – Economics Series* 30. (3): 59–76. Doi: 10.2478/sues-2020-0017
- LADOS, G. and HEGEDŰS, G. 2016. Returning home: An evaluation of Hungarian return migration. *Hungarian Geographical Bulletin* 65. (4): 321–330. Doi: 10.15201/hungeobull.65.4.2
- MAZA, A. 2022. Regional differences in Okun's law and explanatory factors: Some insights from Europe. *International Regional Science Review* 45. (5): 555–580. Doi: 10.1177/01600176221082309
- MELGUIZO, C. 2017. An analysis of Okun's law for the Spanish provinces. *Review of Regional Research* 37. (1): 59–90. Doi: 10.1007/s10037-016-0110-7
- MONTERO-KUSCEVIC, C.M. 2011. *Spatial Features of Okun's Law Using U.S. Data*. Morgantown, WV, West Virginia University Libraries. Doi: 10.33915/etd.3386
- MORAN, P.A.P. 1948. The interpretation of statistical maps. *Journal of the Royal Statistical Society: Series B (Methodological)* 10. (2): 243–251. Doi: 10.1111/j.2517-6161.1948.tb00012.x
- OKUN, A.M. 1963. *Potential GNP: Its Measurement and Significance*. Cowles Foundation Paper 190. Cowles Foundation, New Haven, CT, Yale University.
- OVERMAN, H.G., PUGA, D. and VANDERBUSSCHE, H. 2002. Unemployment clusters across Europe's regions and countries. *Economic Policy* 17. (34): 115–147.
- PALOMBI, S., PERMAN, R. and TAVÉRA, C. 2017. Commuting effects in Okun's law among British areas: Evidence from spatial panel econometrics. *Papers in Regional Science* 96. (1): 191–209. Doi: 10.1111/pirs.12166
- PATACCHINI, E. and ZENOU, Y. 2007. Spatial dependence in local unemployment rates. *Journal of Economic Geography* 7. (2): 169–191. Doi: 10.1093/jeg/lbm001

- PATUELLI, R., SCHANNE, N., GRIFFITH, D.A. and NIJKAMP, P. 2012. Persistence of regional unemployment: Application of a spatial filtering approach to local labor markets in Germany. *Journal of Regional Science* 52. (2): 300–323. Doi: 10.1111/j.1467-9787.2012.00759.x
- PERREIRA, R.M. 2013. *Okun's Law and Regional Spillovers: Evidence from Virginia Metropolitan Statistical Areas and the District of Columbia*. College of William & Mary, Department of Economics, Working Paper Number 140. Williamsburg, VA, William & Mary.
- RÍOS, V. 2017. What drives unemployment disparities in European regions? A dynamic spatial panel approach. *Regional Studies* 51. (11): 1599–1611. Doi: 10.1080/00343404.2016.1216094
- SALVATI, L. 2015. Space matters: Reconstructing a local-scale Okun's law for Italy. *International Journal of Latest Trends in Finance & Economic Sciences* 5. (1): 833–840.
- SASONGKO, G., ARTANTI, N., HURUTA, A. and LEE, C.-W. 2020. Reexamination of Okun's law: Empirical analysis from Panel Granger Causality. *Industrija* 48. (4): 63–80. Doi: 10.5937/industrija48-29455
- SCHABENBERGER, O. and GOTWAY, C.A. 2005. *Statistical Methods for Spatial Data Analysis*. Boca Raton, Champion & Hall/CRC.
- SOYLU, Ö.B., ÇAKMAK, İ. and OKUR, F. 2018. Economic growth and unemployment issue: Panel data analysis in Eastern European countries. *Journal of International Studies* 11. (1): 93–107. Doi: 10.14254/2071-8330.2018/11-1/7
- SZULC, E. and JANKIEWICZ, M. 2018. Spatio-temporal modelling of the influence of the number of business entities in selected urban centres on unemployment in the Kujawsko-Pomorskie voivodeship. *Acta Universitatis Lodzianensis. Folia Oeconomica* 4. (337): 21–37. Doi: 10.18778/0208-6018.337.02
- VILLAVERDE, J. and MAZA, A. 2007. Okun's law in the Spanish regions. *Economics Bulletin* 18. (5): 1–11.
- VILLAVERDE, J. and MAZA, A. 2009. The robustness of Okun's law in Spain, 1980–2004. Regional evidence. *Journal of Policy Modeling* 31. (2): 289–297. Doi: 10.1016/j.jpolmod.2008.09.003
- YERDELEN, F. and İÇEN, H. 2019. Heterogeneous multi-dimensional panel data models: Okun's law for NUTS2 level in Europe. In *Selected Topics in Applied Econometrics*. Eds.: ÇAĞLAYAN, E. and KORKMAZ, Ö., Istanbul, Peter Lang, 31–45.